Stock Strength and Rate of Mortality of the Norwegian Spring Spawners

as indicated by Tagging Experiments in Icelandic Waters.

by

Olav Dragesund Norway and

Jakob Jakobsson Iceland

1. Introduction.

Since 1948 large scale herring tagging experiments have been carried out during the Norwegian winter fishery and the Icelandic North coast summer fishery. Reports were published in 1950 and 1952 (Fridriksson and Aasen), giving a detailed account of the methods applied as well as the very encouraging results which illustrated the validity of the methods. Later, Aasen (1958) dealt with the first estimation of stock strength based on the tagging experiments 1948-1954. A new report based on the experiments up to 1960 is being prepared by the present authors. Since the report is not yet ready for publication, it was considered necessary to present its most relevant section for this symposium.

2. Methods and Material.

2.1. General.

The methods used in this paper for calculating the stock strength and the survival rates are based on the theoretical considerations of Aasen (1958), as well as on those of Beverton and Holt (1957) especially as regards estimation of the instantaneous fishing mortality coefficient F.

The equation

 $\frac{y}{S} = \frac{n}{N} \quad \dots \quad (1)$

where y denotes the fishery yield

S the stock present

n the number of recaught fish

N the tagged fish present

is used in the present paper as basic equation for stock strength calculations.

A necessary condition for this basic assumption, i.e. that untagged and tagged fish are caught in the same proportions, is that the tagged herring are randomly distributed in the stock, since it is not given that the boats fish at random.

Since the reduction plants are scattered along the coast and each plant receives herring which mainly come from a particular part of the fishing grounds, it is reasonable to suppose that if the tags are randomly distributed between the reduction plants the tagged herring will be randomly spread in the stock. We will then exclude the possibility that the tagged herring, which mainly consist of old herring, are recaptured only in the beginning of the season when most of the oldest herring are caught, and not mixed up with the younger year-classes, entering the fishing ground later in the season.

By using the mean number of tags per 100.000 hectolitres of reduced herring as the expected number and comparing it to the actual number of returns per 100.000 hectolitres in each reduction plant, it proved possible to carry out χ^2 tests on the returns of the above mentioned experiments during the six years period 1952-1957 inclusive. Throughout this period, tags from the Icelandic North coast experiments in the preceding summer were randomly distributed (0.90 > p > 0.05) between reduction plants. Unfortunately, in the years 1958-1960 the returns were too few for statistical analyses, but we make the assumption that they were randomly distributed also in this last period.

Generally, the returns from the Norwegian experiments did not comply with the basic assumption (page one) so for that and other reasons which will not be discussed here, only the returns from the Icelandic experiments and those carried out by the R/S G.O. Sars in the open ocean proved suitable for stock size analyses.

The method of estimation of each component in equation (1) will now be discussed.

2.2. y&n.

Since the opinion of the present authors is in conformity with Aasen's theoretical discussion of the parameters y and n in equation (1), they wish to refer to his discussion and only state that here they use

where \overline{a} denotes the average number of individuals per hectolitre and c the landing figure in hectolitres.

The calculated number of returns may then be expressed by the equation:

$$n = \frac{r}{e} \times \frac{c}{p} \quad \dots \quad (3)$$

where r is the actual number of returns

e the efficiency of the magnets

c the landing figure, and

p the quantity reduced in plants with known efficiency.

2.3. N.

When considering N, i.e. the number of tagged fish present on the Norwegian winter herring fishing grounds, we must consider the characteristics of the herring tagged in the Icelandic experiments as well as the succession of events which happen to the herring during the period from liberation in July or (infrequently) August off the North and North-east coast of Iceland till they enter the winter herring fishing grounds off the West coast of Norway.

With regard to the characteristics of the herring of the North coast of Iceland, it must be borne in mind that the herring concentrations are varying mixtures of Icelandic and Norwegian herring tribes.

In order to find N_{N} , i.e. the number of tagged herring which will seek the spawning grounds off the West coast of Norway, we must make the assumption that only spring spawners with Norwegian type of scales will do so (see e.g. Fridriksson, 1944 and 1958). By considering scale analyses of samples taken at the time of the tagging and from the catches in the tagging areas, it proved possible to estimate the proportion of the Norwegian type of scales for each liberation of the tagging experiments during the period 1951-1960 (Fridriksson, 1953-1960 and unpublished data). Having thus estimated N_N (the number of tagged herring with Norwegian type of scales) we proceed to consider the succession of events which will reduce N_N before they reach the fishing grounds off the West coast of Norway the following winter season. These losses can be due to (1) effect of tagging and (2) fishing and other causes including natural mortality. Considering these in turn we have:

2.3.1. Effect of tagging.

Although experiments on herring tagged with internal steel tags show very low mortality and shedding of tags due to tagging, it must be borne in mind that the Icelandic experiments were carried out in unsheltered waters under varying circumstances and the herring used for tagging were taken from different catches, and hence the condition of the herring may have varied from liberation to liberation. By considering the total returns in Norway from each experiment there is a significant variation in returns from the various liberations within the same experiment. This difference can either be due to (a) varying tagging or fishing mortality or (b) non random distribution of tags in the Norwegian catches.

Since the returns from any one liberation within an experiment are too few for testing statistically, whether they are randomly distributed in the Norwegian catches, sufficient number of returns from 3 or 4 liberations (giving the highest percentage recaptures) were taken and tested. Having found these returns randomly distributed and thus ruling (b) out, the percentage returns (A) from such "standard liberation" was calculated. Then the effective number of tagged herring (of the Norwegian type) was $N_{Ne} = \frac{B}{A} \quad 100 \quad \dots \quad (4)$

where A is the percent returns of the "standard liberations" and B the total number of returns from a given experiment (see also Anon, 1959).

2.3.2. Fishing mortality and Other causes.

It is clear that during the period from the tagging (July) to the beginning of the Norwegian winter season (January) the number of tagged herring in the stock will be reduced further by fishing and natural mortality. Since only a very little part of this fishery is reduced in reduction plants, the Norwegian winter fishery will be considered as the sole cause of the instantaneous fishing mortality coefficient F and all other fishing included in "other causes" of the instantaneous mortality coefficient X. Before attempting to arrive at estimates of F and X separately, their sum, or rather the rate of survival, will be considered.

Since the Icelandic tagged herring (Table 2) recaptured in Norway generally show a regular series of returns during the period in question the authors wish to refer to Aasen's discussion of the survival rate and denoting it by

$$\Theta_{1} = \frac{N_{2Ne} \times r_{1}^{(3)}}{N_{1Ne} \times r_{2}^{(3)}}$$
 (5)

where \mathbf{Q}_1 is the survival rate of any one year

N_{1Ne} the effective number of tagged herring in that year

 N_{2Ne} the effective number of tagged herring the following year

r $\binom{(3)}{1}$ the number of returns of N_{1Ne} in the third year

 $r_{2}^{(3)}$ the number of returns from N_{2Ne} the following year.

Using analogus denotation the ratio

is constant after the third year. Thus a series of estimations of the annual survival rate ${\it Q}$ for any given year can be calculated.

Denoting the annual survival rate of two successive years by \mathbf{Q}_1 and \mathbf{Q}_2 and using corresponding indices as used in 5 we have:

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by deviding (6) by (5) we get an estimation of Q_2 and similarly to (5) this ratio is constant for any year after the 4th year. Series of estimates can then be calculated for Q_2 , which is independent of the series calculated from (5).

Further using analogus denotations:

$$Q_{1} \times Q_{2} \times Q_{3} = \frac{N_{4Ne} \times r_{1}^{(5)}}{N_{1Ne} \times r_{4}^{(5)}} \qquad \dots \dots \dots (7)$$

and dividing (7) by (6) yet another independent series of estimates can be calculated for Q_3 . Thus for Q_1 one such series of estimates can be calculated, two for Q_2 , three for Q_3 etc. (Tables 3 and 4). Clearly a relatively accurate estimate of Q, and hence (F+ X), can thus be obtained if the tags from any one experiment are returned in sufficient numbers for several years. This method is, however, limited to the total annual mortality rate and does not give direct information about the reduction of the number of the tagged herring from the time of liberation to the beginning of the Norwegian winter season.

Using, however, N_{Ne} (table 1) i.e. the number of effectively tagged Norwegian herring in equation 1 and solving for S i.e.

$$S = \frac{yN_{Ne}}{n}$$

it is clear that estimates thus obtained for the stock strength S are too high and hence any direct calculations of $F^{\frac{1}{2}}$ from $S^{\frac{1}{2}}$ and yield figures c (Table 1) will be too low. Nevertheless if the obtained values of F and S are applied to calculate $X^{\frac{1}{2}}$ and these figures are used to reduce N_{Ne} from the tagging month (July) to the beginning of the Norwegian herring season (January), i.e. six months $(\frac{X^{\frac{1}{2}}}{2})$, a new estimate of the stock size ($S^{\frac{1}{1}}$) is obtained. This estimate of the stock strength, however, is too small, and hence $F^{\frac{3}{2}\frac{1}{2}}$ becomes too high, but both the figures are nearer to the true values than $S^{\frac{1}{2}}$ and $F^{\frac{1}{2}}$ respectively. A new value (X^{11}) is then calculated and hence new $S^{\frac{3}{2}\frac{1}{2}\frac{1}{2}}$ and $F^{\frac{3}{2}\frac{1}{2}}$. The figures for $S^{\frac{3}{2}\frac{1}{2}\frac{1}{2}}$ and $F^{\frac{3}{2}\frac{1}{2}}$. The calculations should continue until successive estimates approach each other.

Table 5 and 6 show the stock size and F and X resp. according to this method. The resulting estimates of N are shown in Table 1.

In order to get another set of estimates of the instantaneous fishing mortality coefficient due to the Norwegian fishery for comparison with whose calculated from the stock size yield data and the total annual mortality the authors wish to refer to Beverton and Holt's (1957) discussion, pp. 184-191, and their resulting formula (14.15)

$$F = \frac{\frac{n_1}{n_1} \cdot 1 \cdot \log \left(\frac{n_1}{n_2}\right)}{N_{\tilde{O}} \cdot \left(1 = \frac{n_2}{n_1}\right)} \quad \dots \quad (8)$$

where n_1 and n_2 denote the number of recaught fish in two successive

years

 N_{O} the initial number of effectively tagged fish $\boldsymbol{\gamma}$ the time interval

In order to use this equation the fishing intensity should be constant in the period dealt with. This is approximately the case during the Norwegian winter herring fishery, when we compare two successive years. The duration of the Norwegian herring season, however, is only two-three months, and the mortality rate due to fishing during the rest of the year will not be included in the estimates of F.

Even if the fishing intensity outside the Norwegian season is varying with time, and also different from the Norwegian one, an attempt has been made to apply equation (8). In order to get a series of estimates of F (Table 8) N_0 has been recalculated N_0^1 , i.e. the tagged herring present in the beginning of each new tagging year, according to the values obtained for \mathbf{Q} (Table 4).

Further estimation of the natural mortality (X_n) can be obtained by plotting the fishing effort in the different years against (F + X) and fitting a straight line to the data. The effort is calculated as:

The number of Norw. purse seiners x days on grounds $x \frac{\text{Total catch}}{\text{Catch Norw.}}$

It should be noted that the estimate of X_n obtained from these effort data is not directly comparable with that obtained indirectly from equation (8) and the "approach method" since there the mortality rate due to all other causes than the Norwegian winter fishery is included in X, whereas in the former case X_n does not include mortality due to Icelandic, Russian and Norwegian (summer) fishery and fishery carried out by other nations.

3. Results.

3.1. General.

Table 1 (second column) shows the total number of tagged herring during the Icelandic North coast summer seasons from 1951-1959 (in-Tagging experiments before 1951 (i.e. in 1948 and 1950) are clusive). excluded because the returns from these experiments were rather few and the tagging technique had by then not reached the same standard as in later years. The table clearly shows how the proportion of Norwegian herring (ξ_1) gradually decreases from over 0.9 at the beginning of the decade to less than 0.3 in the last years. On the other hand during the years of 1951-1957 the tagging survival rate ξ $_2$ was remarkably steady, only varying from 0.72-0.80 with an average of 0.77. Thus the proportional variations in the calculated number of effectively tagged herring (N_{Ne} = N_T $\xi_1 \xi_2$) are mainly due to the great changes in ξ 1 the proportion of the Norwegian type of herring.

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Table 2 shows the actual number of returns, the number of hectolitres reduced in plants with tested magnets as well as the per mille returns per million hectolitres (in brackets).

The table clearly shows that, generally, the number of returns for any given experiment reach a maximum in the first year after the tagging and then gradually decrease as the years go by. The only exception to this is the experiment of 1951, the returns of which reach a maximum in 1953 instead of 1952. As a result of this the survival rate Q_1 , (Table 3) becomes absurdly high and hence its use for calculations of fishing and natural mortality rates are meaningless. Trusting that there have not been great variations in survival rates from 1952 to 1953, the authors use Q_2 for the purpose of calcualting rates of fishing and natural mortality in both 1952 and 1953.

Whereas estimates of Q_1 were obtained from the rations of returns from the experiments 1951-52 equation (5) estimates of Q_2 are obtained by this method as well as ratios of recaptures from the 1952-53 and the experiments according to equations (5), (6) and (7). Thus Q_2 is the mean of 10 estimates. Similarly Q_3 , Q_4 Q_8 (Table 4) are the means of 9-15 estimates derived according to equations (5), (6) and (7). The estimates obtained show a gradual reduction of the survival rate $Q_1 = e^{-} (F + X)$ during the period 1953 to 1958 inclusive - or from 0.77 to 0.54. The survival rate for 1959 Q_8 on the other hand proved to be absurdly high 1.24. With reference to this it should be noted that in 1960 the per mille returns are generally very high. The only exception of relatively high returns in 1960 are those from the 1959 experiment. Since all three estimates (Table 3) of Q_8 are proportional to the ratio between the high per mille returns of the previous experiments and the relatively low returns from the 1959 experiment the values for Q_8 become too high.

For the purpose of estimating the number of tagged herring present in 1959, the calculated survival rate for that year (Table 3 and 4) can not be used, especially because the general tendency is clearly shown to be decreased survival rate during the period in question. The authors therefore consider themselves justified in using the survival rate of the previous year for the calculation of tags present in 1959 rather than omitting that year altogether. It must, however, be borne in mind that only future series of recaptures can show, whether in this case the above treatment of survival rates is the right one. Excluding these irregularities the series of returns from the Icelandic tagging experiments during the Norwegian winter season clearly show the regularity with which the North Coast Herring of Iceland visits the spawning grounds off western Norway.

Using the number of effectively tagged herring N_{Ne} as shown in Table 1 and the annual survival rates shown in Table 4 (with the exception of 1952 and 1959), the number of tagged herring (N_o) present in the beginning of each new tagging year for all the experiments were calculated and presented in Table 5, along with the calculated number of tags recaptured during the Norwegian winter herring season according to equation (3) and values shown in Table 1.

3.2. Stock size.

Using the data presented in Tables 1, 2, 3 and 4 and applying the methods described in section 2, estimates of y, n and N were calculated (equations 2, 3, 4 and the approach method) and inserted in equation 1 which was then solved for S₁, i.e. the stock size. The results of these calculations are shown in Table 6 and Fig. 1. Judging by these estimates the Norwegian tribe was at a peak at the beginning of the period (1952), then it decreases until 1954. In 1955 there is a secondary recovery of the stock but since 1956 there has been a steady decline in the stock size in 1959 being only a quarter of the 1952 estimates. These results are in good agreement with age analyses and other Norwegian stock size investigations.

3.3. Fishing and Natural Mortality.

With reference to the discussion in 2. 3.2. it is clear that the calculations of the estimates of the instantaneous fishing mortality coefficient (F) due to the Norwegian winter herring fishery and that of all other causes (X) are interrelated and based on the same principles as the stock size calculations and hence the data used for calculations of F and X according to the approach method are the same as used in 3.2. The results of these calculations are presented in Table 6. In order to get another estimate of F and hence X, the data in Table 5 were used to calculate a series of estimates of F and X according to Beverton and Holt's method (equation 8). The results of these are presented in Table 8.

Comparing the results of these two methods (Tables 6 and 8) it is clear that both show the same general trend i.e. that in spite of a sharp increase in the total instantaneous mortality coefficients (F + X)during the period in question (see also Tables 3 and 4) there is no such increase shown in the instantaneous fishing mortality coefficient. Fig. 2 shows how the Norwegian winter herring fishery has decreased since 1956. Since this decrease is accompanied by a general decrease of the stock (Fig. 1) great changes in F cannot be expected. The increase in the instantaneous mortality coefficient (F + X) is on the other hand in good agreement with the increase in other fisheries as shown in Fig. 2.

Fig. 3 shows the results of fitting a straight line to corresponding data of the effort converted from the Norwegian purse seiners. The value of X_n (0.232) is the upper limit of the instantaneous natural mortality coefficient since mortality due to other causes (tagging mortality and shedding of tags) is included in the estimate.

Taking the differences between (F + X) and X_n estimates of the total instantaneous fishing mortality coefficient (F_T) is shown in Table 9.

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Table 9.

Year	1952	1953	1954	1955	1956	1957	1958	1959
$\mathbf{F}_{\mathbf{T}}$	0.029	0.029	0.199	0.184	0.422	0.348	0.384	0.384

Since we are partly dealing with a purse seine fishery where availability often is of great importance fluctuations in the total fishing mortality coefficient (F_t) are to be expected, but in spite of this the data presented in Table 9 show the general tendency of increased total instantaneous fishing mortality coefficient since 1952.

Mean number of individuals per mill, hl (mill, number)	1 CG	•	296,176	290, 696	302, 752	342,811	321, 635	334, 419	299,760	285, 259	285,944	
Quantity processed in plants	ex p	2,553	4,114	2,740	5.573	4,358	4,866	3,816	1,063	1,490	0, 986	
, Yield of Norw, fishery in mill hl,	ð	9,548	8, 822	7,205	11,744	10, 381	12,32	8, 555	3, 713	4,477	3.227	
Nr, of recapt, in the Norw, plants	ц	6	53	212	285	118	142	88	19	22	31	
ed Year of recapt, un		1951	1952	1953	1954	1955	1956	1957	1958	1959	1960	
Number of tagged herring present at the beginning of the Norwegian winter season	N		2806	10420	6484	4039	4013	2330	1264	794		
Number of effectively tagged herring	N _{Ne} ^{≓N} T€1€2	1232	3121	11152	7671	4759	5346	2977	1674	1034	2224	
Tagging survival rate	${\cal E}^{}_2$	0, 77	0.73	0.72	0, 80	0.79	0, 77	0,77	0.77	0.77	0. 77	
Proport, of Norw, spring spawners	ε ₁		0.84	0.92	0, 95	0, 69	0, 72	0.46	0, 29	0,24	0,29	
Actual nr, of tagged herring	NT		5076 *	17308	10181	8783	9241	8443	7550	5644	9946	
Year of tagging		1950	1951	1952	1953	1954	1955	1956	1957	1958	1959	

+) 2012 of these tags were tagged in open ocean by R/S "G,O,Sars"

The actual number (r) of returns in reduction plants and the per mille returns per mill, hl, (in brackets).

					Yea	Year of recapture	. .1			
Year of	Number of	1952	1953	1954	1955	, 1956	1957	1958	1959	1960
tagging	effectively tagged				Quantity proc	Quantity processed in plants (mill, hl,)	i (mill,hl,)	, ,	•	,
	herring	4,114	2. 739	5,572	4,373	4,836	3, 811	1, 063	1,490	0, 986
1951	3,121	53(4,13)	49(5,73)	86(4,95)	30(2,20)	31(2,05)	13(1,09)	1(0, 30)	2(0,43)	11(0,32)
1952	11,552		212(6.70)	327(5.08)	110(2,18)	101(1,81)	63(1,43)	12(0,98)	15(0, 87)	4(0, 35)
1953	7,671			275(6, 43)	100(2,98)	112(3.02)	48(1,64)	9(1,10)	9(0, 79)	8(1,06)
1954	4, 759				118(5,67)	83(3,61)	44(2,43)	11(2,17)	11(1,55)	3(0, 64)
1955	5,293					151(5,90)	80(3.97)	15(2, 65)	23(2,91)	- 3(0,57)
1956	2.977						88(7,75)	23(7, 27)	16(3,61)	14(4,77)
1957	1.674							19(10,68)	22(8,82)	10(6,06)
1958	1,034								22(14, 28)	21 (20, 59)
1959	2,224									31(14,13)

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Year of			Y	ear of	tagging				
recapture	195 2	1953	1954	1955	1956	1957	1958	1959	
1953	0, 85								
1954	0.97	0.77							
1955	1,01	0.74	0, 39						1951
1956	1.14	0,68	0.57	0,35					
1957	0, 76	0,67	0,45	0,28	0.14			·	
Mean:	0,95	0.72	0.47	0.32	0.14				
1954		0.79							
1955		0.73	0.38						
1956		0,60	0, 50	0.31					
1957		0,86	0.59	0,36	0,19				1952
1958		0, 88	0.45	0.37	0,14	0.09			
1959		1.10	0.56	0,30	0.24	0.10	0.06		
Mean:		0,83	0.49	0.34	0.19	0.10	0,06		
1955			0.53						
1956			0.84	0.52					
1957			0.68	0.42	0,21				1953
1958			0.51	0.42	0,15	0.10			
1959			0.51	0.27	0.22	0,09	0.06		
1960			-	-	0,22	0.17	0.05	0.07	
Mean:		·	0.61	0.41	0.20	0,12	0.06	0,07	
1956				0.62					
1957				0 .62	0.32				
1958				0.82	0,30	0.20			1954
1959				0.54	0.43	0.18	0.11		
Mean:				0,65	0.35	0,19	0.11		
1957					0,51				
1958					0.36	0.24			1955
1959					0,81	0,33	0.20		
1958						0.68			
1959						0.41	0.25		1956
1960						0.78	0.23	0,34	
Mean:						0,62	0,24	0,34	
1959							0,62		1957
1960							0,30	0.43	
Mean:		11 - 11 - 11 - 11 - 11 - 11 - 11 - 11	<u></u>	, ,			0,46	0.43	
1960		ł			· · · · · ·			1.44	1958

Survival rates calculated from the North Coast tagging experiments

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	Q_1	Q_2	Q_3	Q_4	Q_5	Q_6	Q_7	\mathbf{Q}_{8}
	1952	1953	1954	1955	1956	1957	1958	1959
	0,92	0.71	0.74	0,64				
	0,95	0.76	0.65	0,68	0.44			
		0,83	0,60	0,68	0,56	0,58	0,60	
			0,61	0.67	0.49	0.60	0,50	1.17
				0,65	0.54	0.,54	0,58	
					0,56	0,52	0.69	
						0,62	0,39	1.42
							0,46	0.91
								1.44
					, <u>, , , , , , ,</u> ,			
Mean:	0,95	0.78	0.62	0,66	0.52	0.56	0.54	1.24

Table 4

The Annual Survival Rates (Q) in the Period 1952 - 1959

The calculated number of tagged herring (N_0) present in the beginning of each new tagging year

and the number of recaptures (n) during the Norwegian winter herring fishing

Year of recapture

	:		1952	1953	3	1954	4	19	1955	1956	<u> </u>	19	1957	1958	58	1959	6	1960
Year of No tagging	z°	ц	No	п	No	ц	°N	u	N	п	No	u	N	п	z°	н	z ^o	ц
	3121	113	2403	129	1850	181	1203	71	794	79	413	29	231	ŝ	125	6	68	3 G
			11552	558	8895		5.782	261	3816	257	1984	141	1111	42	600	45	324	13
					7671	580	4986	237	3291	284	1711	108	958	31	517	27	279	26
1954							4759	280	3141	211	1633	66	915	38	494	33	267	10
1955									5293	384	2752	1.79	1541	52	832	69	449	10
											2977	197	1667	80	006	48	486	46
										,			1674	66	904	66	488	33
1958															1034	66	558	69
1959																	2224	101

Table 5

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1960		4
1959	54	15,403,986
1958	71	21, 282, 960
1957	101	33, 776, 319
1956	<u>Millahla</u> 129	Mill, number, 41,490,915
1955	149	51.078.839
1954	131	39,660,512
1953	134	38,953,264
1952	218	64, 566, 368

	Es	timates o	jf F ^{₽₽₽}	Estimates of $F^{\phi\phi}$ and $X^{\phi\phi\phi}$ estimated from the "approach-method".	estimate	ed from	the "appı	roach-me	thod".
	1952	1953	1954	1955 1956 1957	1956	1957	1958	1959	1960
П ф ф ф	0,041	0,053	0,095	0,073 0	0,101	0, 085	0,101 0,085 0,054	0, 083	r
XPPC	0, 220	0,220 0,208	0, 336	0.336 0.343 0.553 0.495 0.562 0.533	0,553	0, 495	0,562	0,533	Ť

Table 6

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Estimates of F according to equation (14,15) (Beverton and Holt, 1957)

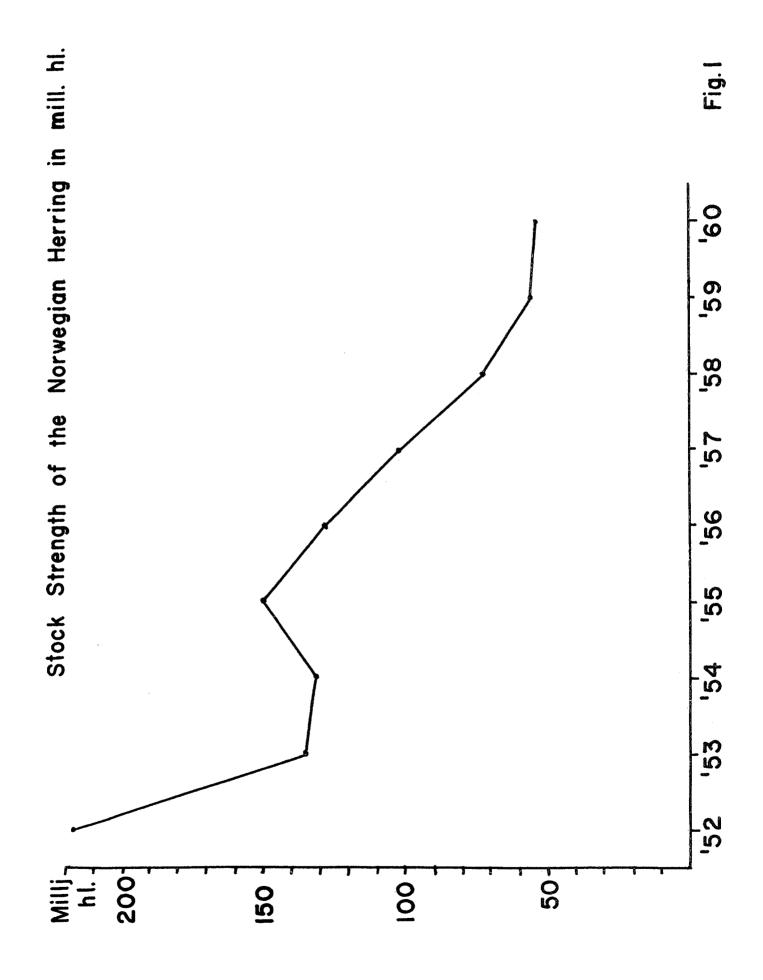
			Year of recapture	ture				
1952	1953	1954	1955	1956	1957	1958	1959	1960
0,034	0,045	0,152	0,057	0,158		ı	4 8	
	0,044	0,121	0,047	060 0	0,123	0,037	1 1	
		0.114	0.044	0, 135	0,111	0,035	0,053	
			0,068	0,096	0,094	0,042	\$	
				0,104	0,113	0,029	I,	
					0,100	0,061	0,054	
						·	0,101	
							0,064	
							ž	
0,034	0,045	0,129	0,054	0,117	0, 108	0,041	0, 068	
	Estima	Estimates of X as the difference between		(F+X), calculated from \mathbb{Q} and F	rom Q and F			
0,227	0,216	0, 302	0,362	0.537	0,472	0.575	0,548	

Table 8

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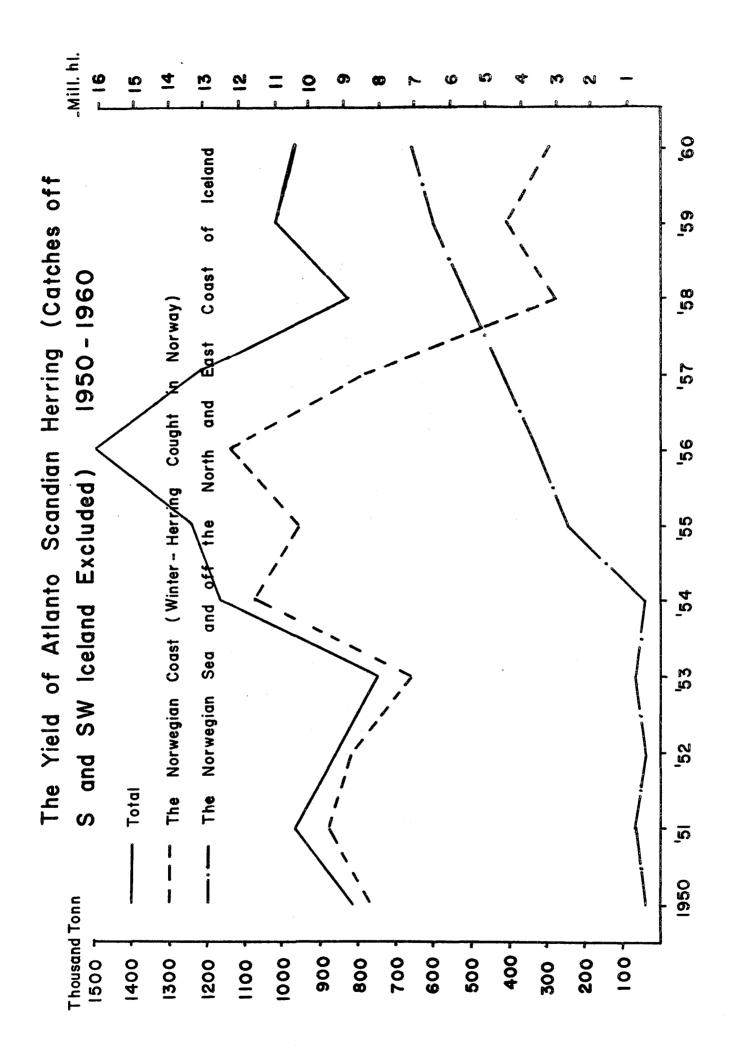


Fig. 2

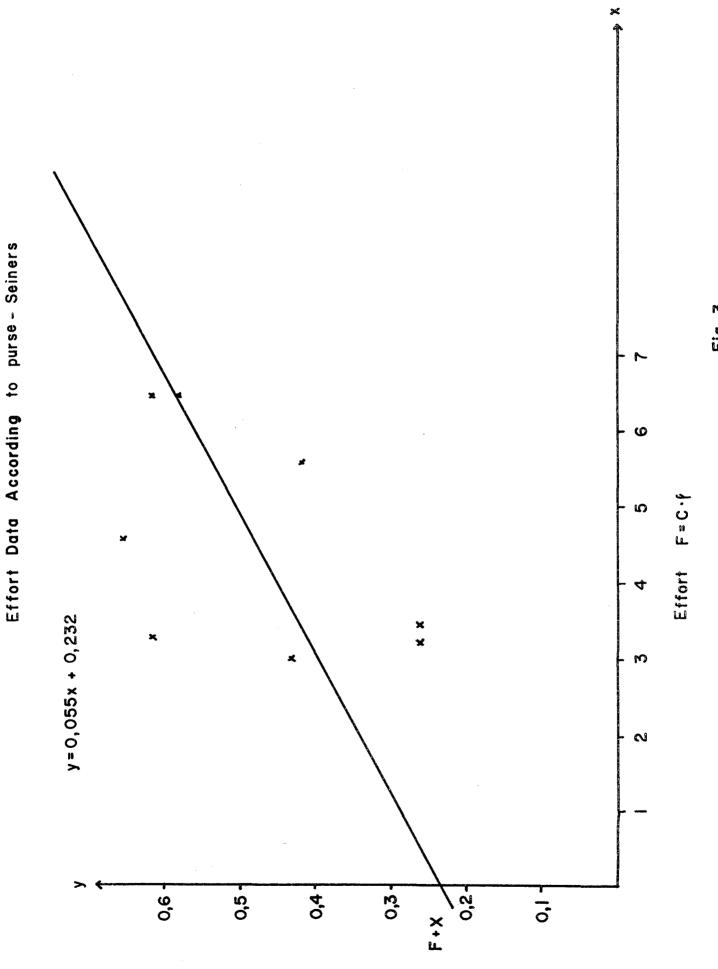


Fig. 3