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# **Does size matter? A bioeconomic perspective on optimal harvesting when price is size-dependent**

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## Abstract

Body size is a key parameter influencing demographic characteristics of fish populations as well as market value of landed catch. Yet in bioeconomic modelling body size is often an overlooked biological and economic parameter. Here we evaluate how size-dependent pricing influences optimal harvest strategies in a model parameterized for two pelagic fisheries, those targeting Atlantic herring (*Clupea harengus*) and Atlantic mackerel (*Scomber scombrus*) in Norway. In our model, positively size-dependent pricing clearly shifts optimal harvest strategies towards lower harvest rates and higher mean body size of caught fish. The results are relatively insensitive to biological (e.g., natural mortality) and economic details of the model (e.g., discount rate or demand function). These findings show that size-dependent pricing influences optimal harvest strategies aiming at maximum economic yield, and hence, require more attention in resource economics and in fisheries management.

Keywords: Bioeconomic modelling, fisheries management, optimal harvest strategies, size-dependent pricing

## 1 **Introduction**

2           It is common that price of fish depends on their size; typically, large individuals fetch a  
3 higher price per kilogram than small individuals of the same species. When present, such size  
4 dependence should be considered when evaluating harvest strategies (Hilborn and Walters  
5 1992). While size-dependent pricing is occasionally included in bioeconomic models, it is very  
6 rare that the consequences of size-dependent pricing per se have been studied in any detail. We  
7 are aware of only few studies where this question has been touched: Gallagher et al. (2004) and  
8 Holland et al. (2005) showed that for Oregon ocean shrimp (*Pandalus jordani*) and rock lobster  
9 (*Jasus edwardsii*), respectively, size-related pricing indeed influences management strategies.  
10 Similarly, Tahvonen (2009) demonstrated in a generic age-structured model how equilibrium  
11 revenue and stock size are affected by size-dependent pricing. More typically, size-dependent  
12 pricing has been either overlooked, acknowledged but not analyzed (Anderson 1989), or taken  
13 as an extrinsically determined model component that is kept fixed in the analysis (e.g., Thunberg  
14 et al. 1998, De Leo and Gatto 2001, Katsukawa 2005). Some studies also acknowledge other  
15 biological factors influencing price, e.g., seasonal changes in fish quality (Larkin and Sylvia  
16 1999). Size-dependent pricing is also important in aquaculture (Bjørndal 1988, Asche and  
17 Guttormsen 2001). Thus, we have all reasons to expect that the size dependency of the price is  
18 widespread and highly relevant for optimal utilization of fish stocks.

19           Here our aim is to quantify how size-dependent pricing influences optimal harvest  
20 strategies. We use Atlantic herring (*Clupea harengus*) and Atlantic mackerel (*Scomber*  
21 *scombrus*) as case studies. Specifically, we focus on the Norwegian spring spawning herring  
22 stock and the Northeast Atlantic mackerel stock, following the stock definitions used in the  
23 management (ICES 2010). Herring and mackerel are of major importance for the fisheries of

24 several European countries, with a total catch of respectively about 1.7 and 0.7 million tonnes in  
25 2009 (ICES 2010). Moreover, herring and mackerel are key species in the northeast Atlantic  
26 pelagic ecosystem (Skjoldal et al. 2004), making their sustainable management even more  
27 important.

28 We utilize an age-structured, discrete-time population model with size-dependent  
29 harvesting. Price data are based on Norwegian market data for herring and mackerel. For  
30 simplicity we assume that the relationship between size and price is linear; this allows us to  
31 smoothly vary the strength of size dependency, in contrast to the earlier studies that only  
32 considered a limited set of fixed pricing scenarios (Gallagher et al. 2004, Holland et al. 2005)  
33 The cost of harvesting is output-regulated and depends on yield. In the optimization the main  
34 target is to maximize the net present value over a long time-scale with discount rates on an  
35 efficient market level, while high discount factors are used to emulate an open-access situation.  
36 We show how optimal fishing mortality and resulting mean individual weight in the catch  
37 depends on the size dependency of pricing. We also estimate the shadow cost, i.e., the  
38 opportunity costs for applying a suboptimal harvest regime and the lost margin of benefit due to  
39 ignoring the size dependency of pricing.

## 40 **Methods**

### 41 **Biological model**

42 The biological model is based on an age-structured population dynamics model with  
43 annual time steps. Sexes are combined because male and female life histories are similar in  
44 herring and mackerel. The parameters and their values are listed in Table A1 (Appendix).

45 There are  $n$  discrete age classes denoted by  $N_a$  with  $a = 1, 2, \dots, n$ , and there is no  
 46 senescence. The last age class is so-called plus-group representing all fish  $n$  years of age or  
 47 older. Population dynamics is then described by

48

$$49 \quad N_a(t+1) = \begin{cases} r(SSB(t)) & \text{if } a = 1 \\ s_a N_{a-1}(t) & \text{if } 1 < a < n, \\ s_n (N_{a-1}(t) + N_a(t)) & \text{if } a = n \end{cases}$$

50

51 where  $SSB$  is the spawning stock biomass,  $r(SSB)$  the stock-recruitment function and  $s_a$  is age-  
 52 specific survival probability defined below.

53 The connection between age and size is defined through the von Bertalanffy growth  
 54 model  $l(a) = L_\infty(1 - e^{-ka})$  where  $L_\infty$  is the asymptotic length and  $k$  the growth coefficient.  
 55 Furthermore, we assume an allometric weight-length relationship  $w(l) = vl^u$ , where  $v$  is the  
 56 length-weight coefficient and  $u$  is the allometric exponent.

57 Reaching maturity is based on the age of the individual and defined by the age-maturity  
 58 ogive  $o(a) = \left(1 + e^{-(a-a_{50})/a_{width}}\right)^{-1}$  with  $a_{50}$  is the age where 50% of the individuals have reached  
 59 maturity and  $a_{width}$  is the coefficient describing width of the maturity envelope. Mature  
 60 individuals constitute the spawning stock, and spawning stock biomass is defined as

61  $SSB = \sum_a w_a o_a N_a(t)$ . For herring recruitment follows the Beverton-Holt stock-recruitment

62 model as defined by (Fiksen and Slotte 2002),  $r(SSB) = \frac{b_1 SSB}{b_2 + SSB}$ , where  $b_1$  is asymptotic

63 maximum recruitment and  $b_2$  determines how fast the asymptote is approached. For mackerel

64 we use the hockey-stick recruitment model estimated by ICES (2010) implying that below a

65 threshold  $SSB_{\max}$  recruitment is a linear, increasing function of  $SSB$ ,  $r(SSB) = b_3 SSB$ , whereas  
66 above the threshold recruitment is assumed to be constant,  $r(SSB) = r_{\max}$ .

67 The fishing mortality is assumed to be length-dependent. Catchability usually increases  
68 with size until levelling off at some intermediate size, a relationship usually described by a  
69 sigmoid curve. However, in our study stocks, no such levelling-off can be detected between  
70 fishing mortality (mean fishing mortality at age from 1984–2009; ICES 2010) and length-at-age  
71 (from the growth model described above): for mackerel, the three-parameter sigmoid model fails  
72 to converge, whereas for herring the inflection point in the resulting fit is far outside the range of  
73 observed lengths. Within the range of observed lengths, a two-parameter power function and the  
74 sigmoid function predict similar fishing mortalities for mackerel, but the power function gives a  
75 lower AIC (difference 2.0) than the sigmoid function. The relationship between fishing mortality  
76 and length is therefore most parsimoniously described as a power function for both stocks. Here  
77 we normalize length relative to mean catch length  $\bar{l}$ , such that

78

$$79 \quad F_a = F(l(a)) = F_0 \left( l / \bar{l} \right)^\eta, \quad (1)$$

80

81 where  $F_0$  is fishing mortality at mean catch length and  $\eta$  is a selectivity parameter that  
82 determines the strength of length dependency; the function is convex for  $\eta > 1$ . The equation  
83 was fitted to fishing mortality at age data from (ICES 2010), transforming age into length using  
84 the von Bertalanffy model described above. Mean catch length  $\bar{l}$  used in the normalization was  
85 estimated through back calculation of mean catch weight based on weight at age in the ICES  
86 catch from 2009 (ICES 2010).



87 The natural mortality  $M$  is based on the values used in the stock assessment (ICES  
 88 2010). The survival probability  $s$  is determined by the total mortality  $Z$ , which is the sum of  
 89 fishing mortality  $F$  and natural mortality  $M$ :  $s_a = e^{-Z_a} = e^{-(F_a+M_a)}$

90 The link between the biological and economic part of the model is the catch equation,  
 91  $H_a = N_a(1 - e^{-Z_a})F_a/Z_a$ , where  $H_a$  is catch numbers at age. Total catch in terms of biomass, or  
 92 yield, is catch numbers times the mean age-specific individual weight,  $Y = \sum_a H_a w_a$ .

### 93 **Economics**

94 We assume that the relationship between size and price can be expressed as a simple  
 95 linear function of weight, such that we can easily change the strength of size dependence. To  
 96 make the parameters easily interpretable, we standardize weight relative to the mean observed  
 97 individual catch weight over all age classes,  $\bar{w}$ . Furthermore, we standardize the price-weight  
 98 coefficient relative to the observed coefficient,  $\beta_0$ . The price function is then

99  $p(w) = p_{\bar{w}} + \beta\beta_0\left(\frac{w-\bar{w}}{\bar{w}}\right)$ , where the intercept  $p_{\bar{w}}$  is the price per unit biomass for  $\bar{w}$ , the

100 observed slope  $\beta_0$  gives the price increase when individual weight is increased by  $\bar{w}$ , and  $\beta$  is  
 101 the relative deviation from the observed slope.  $\bar{w}$  thus acts as the pivot point in this function.

102 The parameters  $p_{\bar{w}}$  and  $\beta_0$  of the price function (Figure 1) were estimated with linear  
 103 regression using data provided by the Norwegian Fisheries Directorate (Per Sandberg, personal  
 104 communication). The data comprise of annual average prices per weight class in 2000–2010. To  
 105 obtain mean weight for each weight class, we assumed that weight at age in the catch is  
 106 normally distributed with mean taken from ICES (2010) and coefficient of variation of 30%. We  
 107 then multiplied the age-specific weight distributions with catch numbers at age (ICES 2010) to

108 obtain total weight distribution in a given year, which allows estimating the mean weight for a  
109 certain weight class as well as the overall mean weight  $\bar{w}$ . Mean price  $p_{\bar{w}}$  was derived from  
110 the data from the Norwegian Fisheries Directorate.

111 We also considered constraining the influence of size on price. As the minimum price we  
112 used the minimum price set by the Norwegian pelagic fish sales organisation, which was 0.60  
113 NOK•kg<sup>-1</sup> for both herring and mackerel in 2010 (Norges Sildesalgslaget 2010). As the  
114 maximum we used the highest mean price in 2010, which was 4.16 NOK•kg<sup>-1</sup> for herring and  
115 11.7 NOK•kg<sup>-1</sup> for mackerel.

116 The total revenue is the sum of the annual yield times the weight-dependent price for  
117 each age class,  $R_t = \sum_a p(w_a)Y_a$ . The cost function is derived from the model of Touzeau et al.  
118 (2000). To avoid potential complications due to the nonlinearity of this cost curve, a simplified  
119 cost function with a linear relationship between yield and cost based on a linear regression of the  
120 cost function of Touzeau et al. (2000),  $C_t = C_0 + \nu Y_t$ , has been used. Here the intercept  $C_0$   
121 represents fixed costs and the term  $\nu Y_t$  variable costs. The net revenue is then  $\Pi_t = R_t - C_t$ .  
122 With  $d$  denoting discount rate, the net present value is the sum of annual discounted net  
123 revenues:

124

$$125 \quad NPV = \sum_0^t \frac{\Pi_t}{(1+d)^t}.$$

126

127 We tested a demand function estimated from Norwegian yield and price data in 2000–  
128 2009. However, the resulting fit was poor because the lack of information about willingness to  
129 pay outside the narrow range of actual production volumes and because of other factors

130 influencing the price. Sensitivity analysis showed no significant influence on the results, and  
131 therefore demand effects were not considered further in this study. However, because of this,  
132 and the simple cost function used, we emphasize that we do not expect the model to give precise  
133 quantitative predictions, even though the results shown are qualitatively robust.

## 134 **Simulations**

135 The model was run over 1000 years, an essentially infinite time horizon unless discount  
136 rate is very low ( $d < 0.01$ ). The first 200 years were used to establish pre-fishing equilibrium.  
137 Fishing started from the virgin stock at year 200, causing the stock to decline towards a new  
138 equilibrium.. It is assumed that size selectivity of the fisheries is fixed, but that the managers can  
139 adjust the overall level of fishing mortality such that net present value (NPV) of the stock is  
140 maximized. Harvest strategies in our model are therefore defined through the parameter  $F_0$ ,  
141 fishing mortality at mean catch length. The optimal  $F_0$  was identified using the function  
142 “optimize” in R (R Development Core Team 2010).

## 143 **Results**

144 The effect of size-dependent pricing is to shift harvest maximizing net present value  
145 (NPV) towards lower values (Figure 2). Comparing the harvest strategy optimized for size  
146 dependency of the current pricing regime and one without any size dependency, the shadow cost  
147 of omitting the size dependency in the current pricing regime is found to be about 3.1% for  
148 herring and 3.8% for mackerel relative to the optimal NPV (discount rate  $d = 0.05$ ). Shadow  
149 costs define here the difference in NPV between the optimal harvest strategy with the current,  
150 size-dependent price regime and the one that would be optimal when we assume no size-  
151 dependent pricing, i.e., the opportunity cost of ignoring size dependency in price.

152           With increasing effect of size on price (increasing  $\beta$ ), the mean fishing mortality that  
153 maximizes NPV decreases and the corresponding mean individual size increases (Figure 3); for  
154 herring the effect is almost linear within the considered range whereas for mackerel the effect is  
155 levelling off for strong size dependence. Mean fishing mortalities (age groups 4-8 for mackerel  
156 and 5-14 weighted by stock numbers for herring) yielding maximum NPV (with discount rate  
157  $d = 0.05$ ) are  $0.146 \text{ year}^{-1}$  in the current pricing regime ( $\beta = 1$ ) compared to  $0.189 \text{ year}^{-1}$   
158 without size-dependent pricing for herring and  $0.242 \text{ year}^{-1}$  compared to  $0.357 \text{ year}^{-1}$  for  
159 mackerel. For reference, the latest stock assessments estimated fishing mortalities at  $0.154 \text{ year}^{-1}$   
160 for herring and  $0.233 \text{ year}^{-1}$  for mackerel (ICES 2010). In our model, the decrease in mean  
161 fishing mortality between no size dependency and a realistic size-dependent pricing is about  
162 23% for herring and 32% for mackerel, and the mean catch weight increases by about 6% and  
163 8%, respectively. The patterns remain similar if the initial stock states are changed from the  
164 pristine levels to lower population abundances estimated in the latest stock assessments (ICES  
165 2010).

166           The discount rate has no influence on the qualitative effect that size-dependent pricing  
167 has on optimal fishing mortality and the corresponding mean size (Figure 3), even when very  
168 high discount rates emulating an open-access situation are considered (Figure 4). Nevertheless,  
169 the discount rate has a considerable quantitative effect on the optimal fishing mortality: the  
170 optimal value increases almost threefold between the extreme cases of no discounting and a very  
171 high discount rate of one, mimicking an open-access situation. As theory suggests, the biggest  
172 effect on optimal fishing mortality is found on low to intermediate discounting levels.

173           Because natural mortality is an important parameter in age-structured models, we  
174 investigate its influence further. For both species, changing natural mortality results in changes

175 in optimal fishing mortality and, to somewhat lesser extent, in the resulting mean catch weight  
176 (Figure 5). These changes are opposite for the two species: higher natural mortality leads to  
177 higher optimal fishing mortality for mackerel but lower optimal fishing mortality for herring;  
178 these effects are slightly more pronounced for low  $\beta$ : an increasing  $\beta$  tends to dampen the  
179 influence of changing natural mortality. The qualitative difference between the species is caused  
180 by the age-dependent natural mortality in herring: assuming a constant natural mortality for  
181 herring leads to mackerel-like results. For catch weights, the influence of changing natural  
182 mortality is qualitatively the same for both species: mean catch weight is increased (decreased)  
183 by higher (lower) natural mortality.

184 Comparing the continuous price function with the empirical step functions either in 2009  
185 or 2010 given in Figure 1 shows that the differences are negligible: optimization with discrete  
186 price classes displays the same shift between constant price and size dependency and similar  
187 quantitative results. The same applies when applying minimum and/or maximum price caps.  
188 Additionally, optimal fishing mortality depends on the selectivity parameter  $\eta$ , but the effect is  
189 negligible (mackerel) or does not influence the qualitative pattern (herring) (Figure 6).

## 190 **Discussion**

191 The size of fish is a key parameter from biological as well as technical and economic  
192 perspectives. In terms of biology, size is closely linked with growth, maturation, reproductive  
193 output and survival. Avoidance of growth overfishing — catching fish too small, before a cohort  
194 has realized its growth potential — has been a crucial part of fisheries management since  
195 Beverton and Holt (1957). Using minimum size restrictions and mesh size regulations as  
196 management tools to protect the productivity of a fish stock reflects this idea. Likewise, the  
197 fishing industry has a strong interest in fish size both to optimize the industrial utilization and to

198 serve consumer preferences. Ultimately, these market incentives are a key driver for size-  
199 dependent pricing. Yet the connection between price and body size is an understudied issue in  
200 fisheries economics. This is surprising given that size-dependent pricing is ubiquitous and that  
201 its role in rational harvest management has in theory been acknowledged for a long time  
202 (Hilborn and Walters 1992). While a number of studies on capture fisheries have included size-  
203 dependent pricing, the majority of these have only considered a single scenario, without  
204 studying the effects of size-dependent pricing per se (e.g., Helser et al. 1996, Thunberg et al  
205 1998, Katsukawa 2005). A few studies have included a small number of alternative size-  
206 dependent pricing scenarios when analysing specific fisheries (Gallagher et al. 2005, Holland et  
207 al. 2005). There are also a few generic models that have included size-dependent pricing  
208 (Anderson 1989, Tahvonen 2009). While Anderson (1989) did not elaborate on the importance  
209 of size-dependent pricing, Tahvonen (2009) showed with a concrete example how positively (or  
210 negatively) size-dependent pricing leads to equilibrium revenues to be maximized at a higher  
211 (respectively lower) stock biomass level. The topic has also been discussed outside fisheries  
212 economics. For example, price of fish produced in aquaculture is size-dependent, and this has  
213 consequences for optimizing the production cycle (Bjørndal 1988, Asche and Guttormsen 2001).

214 Our goal here was to demonstrate the importance of size-dependent pricing for  
215 economically optimal harvest strategies using simple age-structured models parameterized for  
216 two important pelagic fish stocks. The results confirm our expectations: fishing pressure  
217 maximizing economic yield in terms of revenue and net present value depends on the applied  
218 size-based pricing regime, with implications for mean size of fish in the catches and equilibrium  
219 stock abundance. The results reveal potential for overestimation of future profits and rent  
220 dissipation due to the application of suboptimal harvest strategies when ignoring size-dependent  
221 pricing; the effect in the examples studied here is nevertheless quantitatively modest. However,

222 the effect on the resultant harvest strategy is appreciable. Considering that in most commercially  
223 important fish stocks the price is size-dependent, the insights presented here call for broader  
224 utilization of size-dependent pricing in economic modelling.

225         When the size-price relationship is monotonic and positive, maximum economic yields  
226 are obtained through higher stock sizes and smaller catches than when such size dependence is  
227 absent. In other words, positive size-price relationships imply that lower fishing mortalities  
228 optimize the net present value. It is acknowledged that catches below MSY typically maximize  
229 the resource rent (Clark 2006, Grafton et al. 2007, Tietenberg and Lewis 2008). Our results  
230 suggest that this effect might even be stronger than MEY estimates ignoring size-dependent  
231 pricing would suggest. Therefore, taking positively size-dependent pricing into account would  
232 result in lower harvest rates, which would not only ultimately increase society's direct economic  
233 benefits from fisheries, but also reduce the negative effects fisheries might have on the  
234 ecosystems. This is in accord with the precautionary approach to fisheries management, calling  
235 for more conservative harvest policies that lead to an increased overall stock abundance. This  
236 would enhance population resilience, harvest productivity and efficiency.

237         Our results resonate with the calls to save the big fish for the sake of improved biological  
238 sustainability (Birkeland and Dayton 2005, Francis et al. 2007, Diekert et al. 2010). This  
239 statement may seem counterintuitive since we emphasize the high market value of large-sized  
240 fish. However, we have focused on a single-owner fishery where the manager adopts a long time  
241 perspective, avoiding the short-sighted temptation of targeting the large fish. In our model the  
242 opportunity cost of choosing a suboptimal harvest strategy can only be mitigated by reducing the  
243 overall fishing mortality, therefore allowing more fish to reach large sizes, and ultimately,  
244 leaving more big fish in the sea.

245           We emphasize that in general, mesh size regulations and other means of modifying size-  
246 dependent catchability can be important levers of optimizing the harvest, even though we did not  
247 consider them here. Our model applies to pelagic fisheries where effective regulation of size-  
248 dependent catchability is difficult; gears like purse seines are weakly size-selective, and pelagic  
249 fish have low survival after slipping through meshes or being strangled in a net. However,  
250 demersal fish are often more robust and have better chances of surviving if slipping through  
251 meshes or sorting grids, or discarded after the capture. Mesh size regulations can therefore be a  
252 means to ensure that enough fish have chance to grow to most valuable size classes. In such  
253 cases, economic implications of considering size-dependent pricing are likely larger than what  
254 we have found for herring and mackerel here.

255           We have used the simplest possible price function, a linear relationship between price  
256 and weight, as the default price function. While this is a good approximation for herring and  
257 mackerel, more complex relationships occur in some fisheries (Thunberg et al. 1998). There is  
258 also some evidence that the price premium for larger size disappears for very large mackerel  
259 (Fig. 1), a situation which may not be uncommon if very large fish have lower flesh quality,  
260 higher concentrations of contaminants, or are less suited for industrial processing. Similarly, if  
261 small-sized fish are destined to fish meal production instead of human consumption, their  
262 precise size may become unimportant. We therefore also tried constraining the price between  
263 certain minimum and/or maximum levels, but this did not change the results in any essential  
264 way (the bulk of the harvest comes from medium-sized fish). Additionally, our price function  
265 implies continuously increasing effect of size on price, whereas real markets commonly operate  
266 with discrete weight classes. While using the real weight classes should give more precise  
267 results, our simulations suggest that a linear function is a good approximation. Moreover, market



268 weight classes often vary over time and space. For these reasons we consider a linear  
269 approximation as justified here.

270 Our study has several general limitations. We rely on a single species approach and focus  
271 in the steady state scenario without environmental oscillations. Size structure in the model is  
272 determined by the age structure using a static age-length relationship. In reality, size structure of  
273 the harvestable stock is determined by intra- and inter-specific interactions (availability of  
274 resources, and presence of predators). In particular, density-dependent body growth (e.g.,  
275 Lorenzen and Enberg 2002) might counter the increase in average size that is expected to occur  
276 when fishing is reduced and therefore influence management strategies (Helser and Brodziak  
277 1998). A significant source of uncertainty is natural mortality. This is modelled as constant  
278 (mackerel) or age-dependent (herring); both assumptions are crude approximations of reality but  
279 unfortunately data for more realistic choices are not available. Also our economic model is  
280 rather simplistic, apart from inclusion of size-dependent price. We assume a simple yield-  
281 determined cost function, and absence of market feedbacks on price. Our initial explorations  
282 suggested that considering effort-dependent costs or demand curves is not important for the  
283 questions addressed here. Moreover, developing a more detailed bioeconomic model for  
284 mackerel would have been a major task on its own, whereas for herring, more detailed  
285 information exists (e.g., Touzeau et al. 2000, Sandberg 2006). Finally, the model does not  
286 account for practical challenges in balancing stakeholder interests nor policy implementation  
287 and enforcement; in reality, suboptimal management due to political and social pressures and  
288 illegal fishing are more a rule than an exception (Beddington et al. 2007, Agnew et al. 2009,  
289 Mora et al. 2009).

290 Another angle worth discussing is that of fisheries-induced adaptive changes. Statistical,  
291 experimental and modelling approaches are giving increasing support to the hypothesis that

292 fishing causes evolutionary shifts in life-history traits (Jørgensen et al. 2007, Dunlop et al. 2009,  
293 Sharpe and Hendry 2009). Because these changes typically involve reduced adult body size  
294 (Heino 1998, Enberg et al. 2011), size-dependent pricing is relevant also when trying to estimate  
295 the possible economic impacts of fisheries-induced evolution.

296 To conclude, this study has highlighted that size-dependent pricing has important  
297 implications for fisheries management. Because markets usually value large fish more than  
298 small fish, harvest strategies should consider how the harvest influences size structure of the  
299 catch: the higher the harvest pressure, the lower the mean size of fish. Therefore, harvest  
300 policies that aim at maintaining harvest of large-sized fish lead to lower harvest levels than those  
301 that ignore the quality of the catch; ultimately, this might lead to more sustainable harvest and  
302 increased economic benefits from the fisheries. We encourage further explorations into  
303 consequences of size-dependent pricing on fisheries management, both in the direction of further  
304 generalizations and towards more detailed studies on specific fish stocks and fisheries.

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309

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## Appendix

Table A1: List of parameters and their values.

Parameters	Definition	Units	Values	
			Herring	Mackerel
$n$	maximum age class <sup>2</sup>	year	15	12
$M_a$	natural mortality <sup>2</sup>	year <sup>-1</sup>	$\left\{ \begin{array}{l} 0.9 \\ 0.15 \text{ for } a = \\ 0.5 \end{array} \right\} \left\{ \begin{array}{l} a \leq 2 \\ 2 < a < \\ n \end{array} \right.$	
$b_1$	asymptotic recruitment <sup>1</sup>	-	$1.9 \cdot 10^{11}$	

$b_2$	steepness of recruitment function <sup>1</sup>	kg	$5.1 \cdot 10^9$	
$b_3$	initial slope of recruitment function <sup>2</sup>	$\text{kg}^{-1}$		1.811
$r_{\max}$	maximum recruitment <sup>2</sup>	-		$4.252 \cdot 10^9$
$L_{\infty}$	asymptotic length <sup>3</sup>	mm	370	418
$k$	von Bertalanffy growth parameter <sup>3</sup>	$\text{year}^{-1}$	0.26	0.43
$v$	length-weight coefficient <sup>3</sup>	$\text{kg} \cdot \text{mm}^{-u}$	$2.32 \cdot 10^{-5}$	$3.4 \cdot 10^{-3}$
$u$	length-weight exponent	-	2.81	3.24
$a_{50}$	age at 50% maturity <sup>2</sup>	year	4	3
$a_{\text{width}}$	width of maturity envelope	year	0.2	0.2
$\eta$	selectivity parameter <sup>2</sup>	-	4.3	3.3
$\beta_0$	price function slope <sup>4</sup>	$\text{NOK} \cdot \text{kg}^{-1}$	2.47	8.65
$\bar{w}$	mean catch weight <sup>2,4</sup>	kg	0.295	0.462
$C_0$	fixed costs <sup>5</sup>	NOK	$3.4 \cdot 10^5$	$3.4 \cdot 10^5$
$v$	cost function slope <sup>6</sup>	$\text{NOK} \cdot \text{kg}^{-1}$	0.7	0.83

<sup>1</sup>(Fiksen and Slotte 2002)

<sup>2</sup>(ICES 2010)

<sup>3</sup>(Jennings and Beverton 1991)

<sup>4</sup>(Norwegian Directorate of Fisheries, direct communication)

<sup>5</sup>(Touzeau et al. 2000)

<sup>6</sup>(Froese and Pauly 2009)

Figure 1: Linear weight-price functions (solid lines) estimated from real price data (dots) and step functions based on price per weight categories for herring (a, c) and mackerel (b, d). The actual price slopes  $\beta_0$  are 2.43 (herring) and 8.97 NOK•kg<sup>-1</sup> (mackerel). For reference, also the case with constant price ( $\beta = 0$ ; grey line) is shown. The pivot point of these lines corresponds to the observed mean individual catch weight,  $\bar{w}$  (dotted line), and price per kilogram for this weight,  $p_{\bar{w}}$ . The price data are annual mean prices per weight class from Norwegian spring-spawning herring and Northeast Atlantic mackerel in 2000–2010 provided by the Norwegian Fisheries Directorate (Per Sandberg, personal communication). We use annual catch weight distributions (ICES 2010) to transform weight classes into mean weight of each weight class (a, b). Step functions (c, d) represent prices per weight class in 2009 (dashed line) and 2010 (solid line).

Figure 2: Net present value (NPV) for herring (a) and mackerel (b) fisheries as a function of harvest proportion under the current pricing regime (solid line), no size-dependent pricing ( $\beta = 0$ ; short dashed line) and a twofold price slope ( $\beta = 2$ ; long dashed line). The distance between vertical grey lines illustrates the difference in the NPV-maximizing harvest strategies in the presence and absence of size-dependent pricing. Discount rate  $d = 0.05$ . Harvest proportion is calculated from the mean fishing mortality over reference ages as defined by ICES (ages 5–14 years for herring, 4–8 years for mackerel; ICES 2010).

Figure 3: The influence of relative price slope  $\beta$  on optimal fishing mortality (black) and the resulting mean individual weight (grey) in the catch of herring (a) and mackerel (b) with respect to different discount rates  $d$ . Discount rate takes values  $d = 0.05 \pm 0.05$  ( $d = 0$  dashed,  $d = 0.1$  dotted line). Optimal strategy is the one maximizing the net present value of a pristine stock. Strength of size dependence of price is expressed as the relative price slope

$\beta$ , with  $\beta = 1$  corresponding to the current price dependence shown in Figure 1 and  $\beta = 0$  to no price dependence. Fishing mortality is the mean fishing mortality over reference ages as defined by ICES (ages 5–14 years weighted by stock numbers for herring, 4–8 years for mackerel; ICES 2010).

Figure 4: Mean fishing mortality optimized for net present value in herring (a) and mackerel (b) fisheries under the assumptions of discount rates between 0 and 1. The continuous line shows results for the real size dependence of pricing. The envelope with dashed lines shows the case of no size dependency (short dashed) and doubled slope (long dashed). Optimal strategy is the one maximizing the net present value of a pristine stock. Fishing mortality is the mean fishing mortality over reference ages as defined by ICES (ages 5–14 years weighted by stock numbers for herring, 4–8 years for mackerel; ICES 2010).

Figure 5: Mean fishing mortality (black) that maximizes NPV in herring (a) and mackerel (b) fisheries and the corresponding mean catch weight (grey) as functions of the relative price slope. The lines represent the default natural mortality  $M_a$  (solid) and  $M_a$  decreased (long dashed) or increased (short dashed) by 20%. The discount rate is  $d = 0.05$ . Optimal strategy is the one maximizing the net present value of a pristine stock. Strength of size dependence of price is expressed as the relative price slope  $\beta$ , with  $\beta = 1$  corresponding to the current price dependence shown in Figure 1 and  $\beta = 0$  to no price dependence. Fishing mortality is the mean fishing mortality over reference ages as defined by ICES (ages 5–14 years weighted by stock numbers for herring, 4–8 years for mackerel; ICES 2010).

Figure 6: Mean fishing mortality (black) optimized for NPV in herring (a) and mackerel (b) fisheries and the corresponding mean catch weight (grey) as functions of the relative price slope. The lines represent the default selectivity parameter  $\eta$  (solid line) and  $\eta$  decreased



(long dashed line) or increased (short dashed line) by 20%. The discount rate is  $d = 0.05$ . Optimal strategy is the one maximizing the net present value of a pristine stock. Strength of size dependence of price is expressed as the relative price slope  $\beta$ , with  $\beta = 1$  corresponding to the current price dependence shown in Figure 1 and  $\beta = 0$  to no price dependence. Fishing mortality is the mean fishing mortality over reference ages as defined by ICES (ages 5–14 years weighted by stock numbers for herring, 4–8 years for mackerel; ICES 2010).

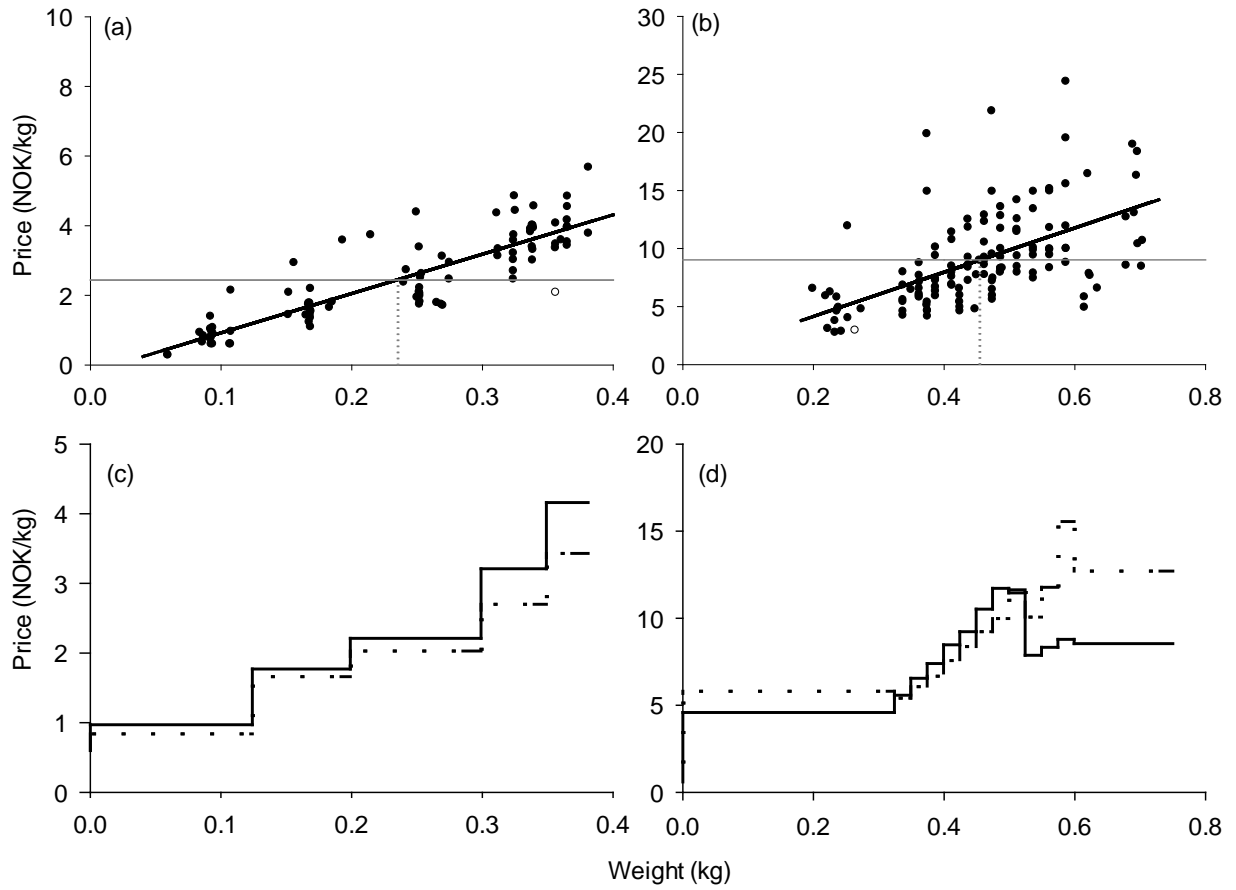


Figure 1

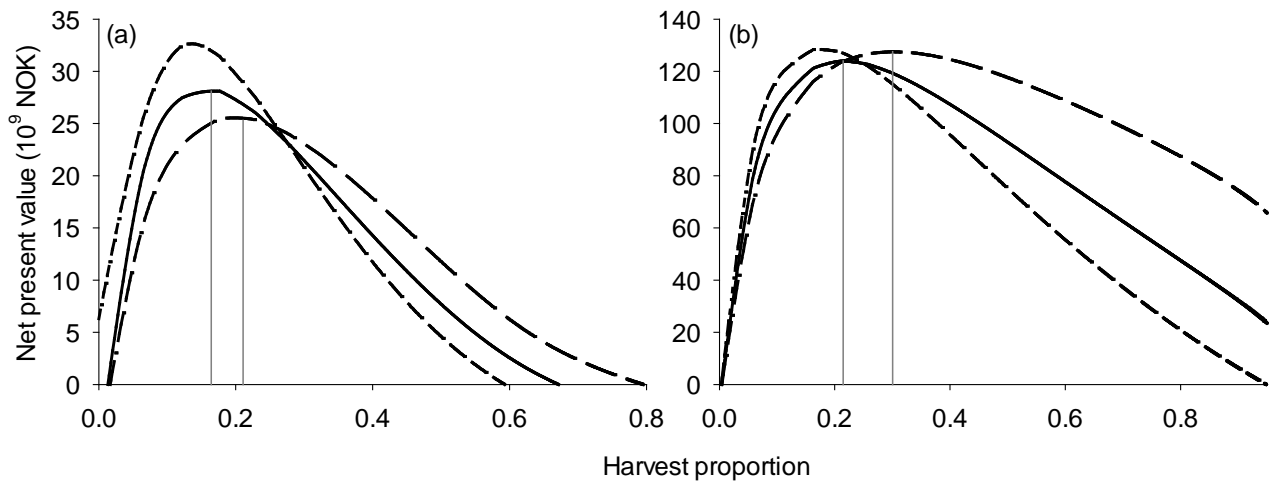


Figure 2

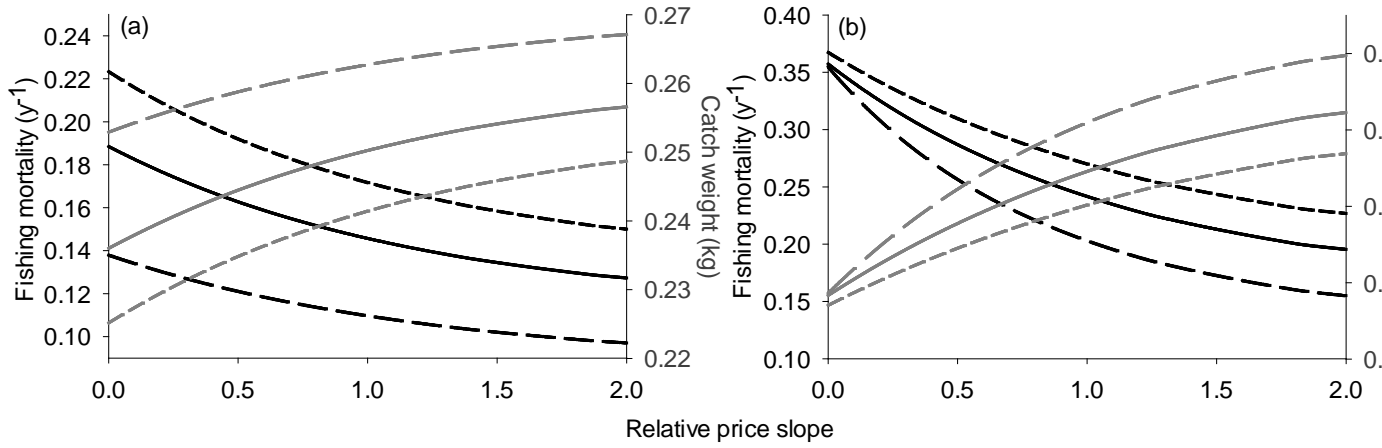


Figure 3

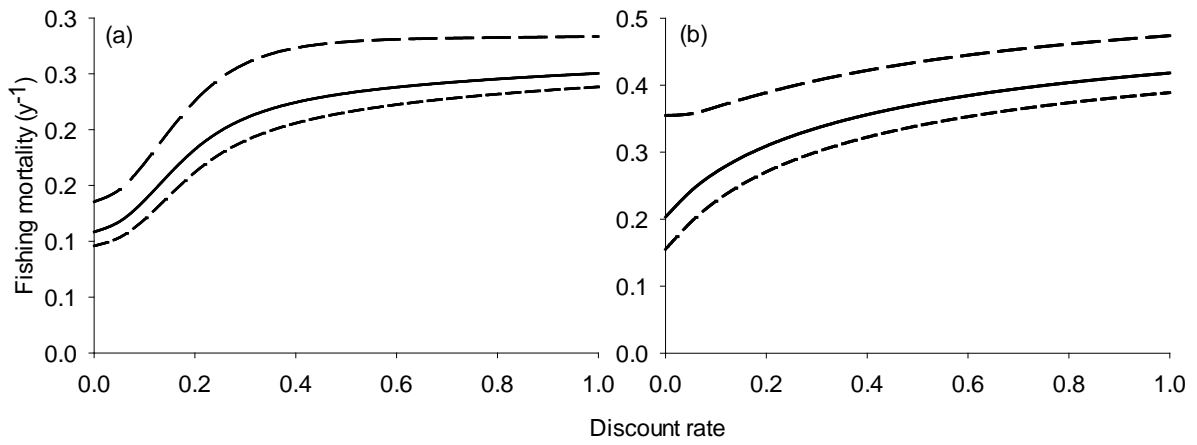


Figure 4

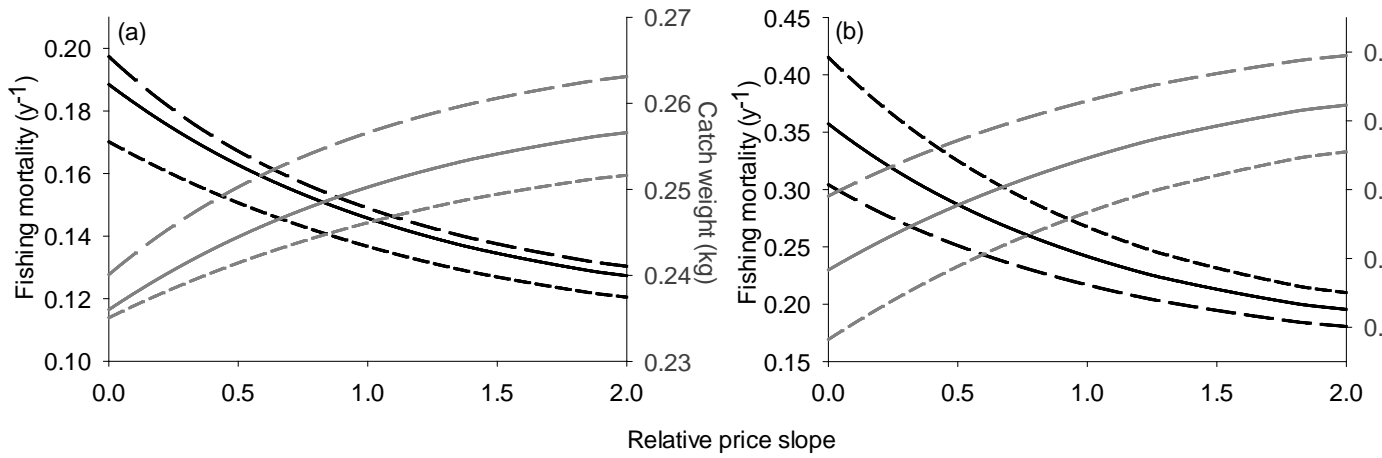


Figure 5

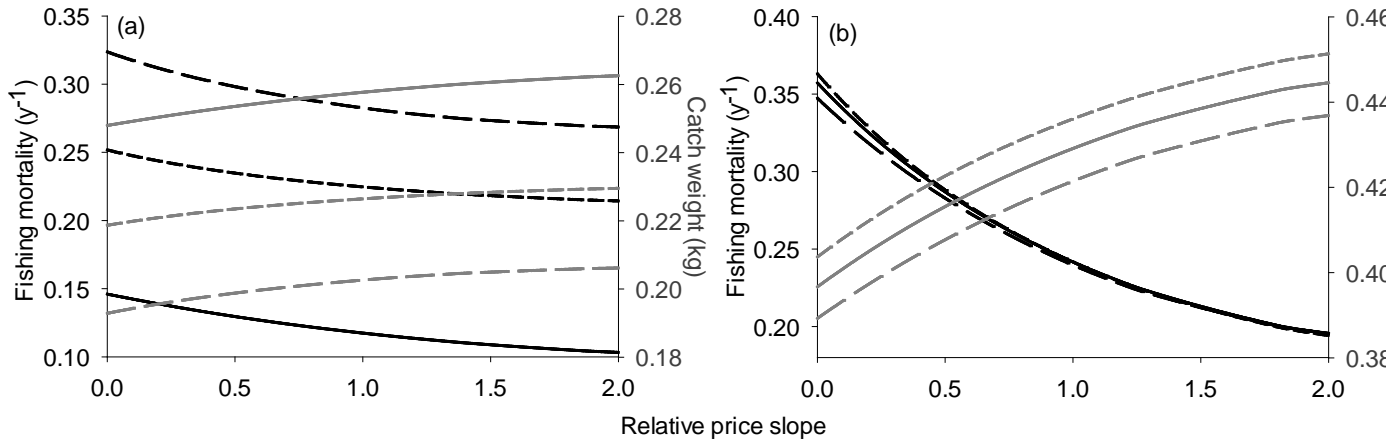


Figure 6